Urns
Game 1
A: all red
B: 50\% red, $50 \%$ blue
Draw red. $\operatorname{Pr}[\mathrm{A}]$ ?
Game 2
A: 70\% red, $30 \%$ blue
B: 30\% red, $70 \%$ blue
Draw $\{3$ red, 1 blue $\} . \operatorname{Pr}[\mathrm{A}]$ ?
Draw $\{10$ red, 7 blue $\} . \operatorname{Pr}[A]$ ?

## Bayes' rule

Hypotheses, data, subjective probability
Hypothesis equivalent to probability distribution over data
Posterior $\propto$ Prior * Likelihood
$\operatorname{Pr}[\mathrm{H} \mid \mathrm{D}]=\operatorname{Pr}[\mathrm{H}] * \operatorname{Pr}[\mathrm{D} \mid \mathrm{H}] / \operatorname{Pr}[\mathrm{D}]$
Simple derivation from joint probability, $\operatorname{Pr}[\mathrm{H}, \mathrm{D}]$
$\operatorname{Pr}[\mathrm{H} \mid \mathrm{D}] \propto \operatorname{Pr}[\mathrm{H}, \mathrm{D}]$
Bernoulli sampling
$x \in\{0,1\}, \operatorname{Pr}[x=1]=q$
Estimate $q \in[0,1]$
Prior density $p(q)$
Likelihood
$\operatorname{Pr}[x=1 \mid q]=q$ (graph)
$\operatorname{Pr}[x=0 \mid q]=1-q$ (graph)
Posterior density
$p(q \mid x)=p(q) \cdot \operatorname{Pr}[x \mid q] / \int p\left(q^{\prime}\right) \operatorname{Pr}\left[x \mid q^{\prime}\right] \mathrm{d} q^{\prime}$
Many observations
$\mathbf{x}=\left(x_{1}, \ldots x_{m+n}\right), \Sigma\left(x_{i}=0\right)=m, \Sigma\left(x_{i}=1\right)=n$
$\operatorname{Pr}(\mathbf{x} \mid q)=(1-q)^{m} q^{n}$
Assume uniform prior, $p(q)=1$
$p(q \mid \mathbf{x})=(1-q)^{m} q^{n} / \int(1-r)^{m} r^{n} \mathrm{~d} r=(1-q)^{m} q^{n} / \mathrm{B}(n, m)$
Beta function: $\mathrm{B}(\alpha, \beta)=\Gamma(\alpha+\beta) / \Gamma(\alpha) / \Gamma(\beta)=(\alpha+\beta-1)!/(\alpha-1)!/(\beta-1)$ !
Beta distribution: $q \mid \mathbf{x} \sim \operatorname{Beta}(n+1, m+1)$
Assume arbitrary beta prior, $q \sim \operatorname{Beta}(\alpha, \beta)$
$p(q \mid \mathbf{x}) \propto(1-q)^{\beta} q^{\alpha} / \mathrm{B}(\alpha, \beta) \cdot(1-q)^{m} q^{n} \propto \mathrm{~B}(\alpha+n, \beta+m)$
simple updating rule: just count 0 s and 1 s
$\alpha, \beta$ "virtual counts"

## Conjugate prior

Given a parameterized family of likelihoods
Bernoulli, $q$
Gaussian, $\mu$
a conjugate family of distributions
Beta, $m$, $n$
Gaussian, $m, s^{2}$
If prior is in conjugate family, posterior is too
Conjugate family closed under multiplication by likelihood
Likelihood viewed as function of parameter being learned $(q, \mu)$
Efficiency of Bayesian updating

## Kalman Filter

Tracking a stochastic process with noisy observation
Generative model
Dynamics: $x_{n}=x_{n-1}+\eta_{n}$
$\eta_{n} \sim \mathcal{N}\left(0, \sigma_{\eta}{ }^{2}\right)$
Gaussian random walk
Observation: $y_{n}=x_{n}+\varepsilon_{n}$ $\varepsilon_{n} \sim \mathcal{N}\left(0, \sigma_{\varepsilon}^{2}\right)$
Gaussian noise
Independence: $\perp\left\{\eta_{n}, \varepsilon_{n} \mid n \in \mathbb{N}\right\}$
Causal graphical model
Conjugate prior
Gaussian likelihood, parameterized by the mean: $y_{n} \sim \mathcal{N}\left(x_{n}, \sigma_{\varepsilon}^{2}\right)$
Gaussian prior, parameterized by mean and variance
$x_{n} \sim \mathcal{N}(a, b)$
Posterior

$$
\begin{aligned}
& p\left(x_{n} \mid y_{n}\right) \propto e^{-\frac{1}{2 b}\left(x_{n}-a\right)^{2}} \cdot e^{-\frac{1}{2 \sigma_{\varepsilon}^{2}}\left(x_{n}-y_{n}\right)^{2}} \\
& \propto e^{-\left[\left(\frac{1}{2 b}+\frac{1}{2 \sigma_{\varepsilon}^{2}}\right) x_{n}^{2}-2\left(\frac{1}{2 b} a+\frac{1}{2 \sigma_{\varepsilon}^{2}} y_{n}\right) x_{n}\right]} \\
& \propto e^{-\left(\frac{1}{2 b}+\frac{1}{2 \sigma_{\varepsilon}^{2}}\right)\left(x_{n}-\frac{\frac{1}{b} a+\frac{1}{\sigma_{\varepsilon}^{2}} y_{n}}{\frac{1}{b}+\frac{1}{\sigma_{\varepsilon}^{2}}}\right)^{2}}
\end{aligned}
$$

Precision-weighted averaging

$$
x_{n} \left\lvert\, y_{n} \sim \mathcal{N}\left(\frac{\frac{1}{b} a+\frac{1}{\sigma_{\varepsilon}^{2}} y_{n}}{\frac{1}{b}+\frac{1}{\sigma_{\varepsilon}^{2}}}, \frac{1}{\frac{1}{b}+\frac{1}{\sigma_{\varepsilon}^{2}}}\right)\right.
$$

Iterative prior
$x_{n} \mid y_{n} \sim \mathcal{N}(c, d)$
$x_{n+1} \mid y_{n}=x_{n}+\eta_{n} \sim \mathcal{N}\left(c, d+\sigma_{\eta}{ }^{2}\right)$
Convolution: $p(\alpha+\beta=X)=\int p(\alpha=Z) p(\beta=X-Z) \mathrm{d} Z$
Update rules
$x_{n} \mid \mathbf{y}_{n-1} \sim \mathcal{N}\left(\mu_{n}, s_{n}^{2}\right)$
$\mu_{n+1}=\frac{\frac{1}{s_{n}^{2}} \mu_{n}+\frac{1}{\sigma_{\varepsilon}^{2}} y_{n}}{\frac{1}{s_{n}^{2}}+\frac{1}{\sigma_{\varepsilon}^{2}}}=\frac{\sigma_{\varepsilon}^{2} \mu_{n}+s_{n}^{2} y_{n}}{\sigma_{\varepsilon}^{2}+s_{n}^{2}}$


## Exercises

1. Change the initial prior in the naive Bernoulli estimator. For example, set it to zero for all $\mathrm{p}<.5$ (and remember to normalize). What do you think will happen? Run the model and see, then explain.
2. Extend the Bernoulli learner to ternary observations. (If you like cover stories: You're training for a roshambo match and watching video of your opponent's past matches to estimate his tendencies.)
a) For binary observations, the latent parameter to be estimated is a single number, $q \in[0,1]$. How would you characterize the thing to be estimated for ternary observations? (Answer before moving on.)
b) There are multiple good answers for part a, but let's go with this: a vector $Q=\left(q_{1}, q_{2}, q_{3}\right)$, constrained to satisfy $\Sigma\left(q_{i}\right)=$ 1. The probability of rock is $q_{1}$, paper $q_{2}$, and scissors $q_{3}$. Assume a uniform prior for $Q$. What's the posterior after a single observation of rock? What's the posterior after observing $k$ rocks, $m$ papers, and $n$ scissors? You can answer without the normalization (using $\propto$ ), or if you like integrals you can work out the normalization constant.
c) Based on part b, you should be able to tell what the conjugate prior is for a ternary random variable. Try to extend this to the conjugate prior for a general n -ary random variable. Write out the distribution (with or without normalization), and write out the update rule for how the distribution's parameters change each time a new observation is made.

The answer to part c is called a Dirichlet distribution. Look it up in Wikipedia (a reliable resource for this sort of thing) and compare the expression there for the probability density (PDF) to your answer above.
d) Download http://matt.colorado.edu/teaching/mathmodeling/plotDirichlet.m for making nice pictures of Dirichlet distributions in the ternary case. (I drew it as an equilateral triangle, with each vertex corresponding to one of the values of the ternary observable variable.) Try different values and make some observations. Try non-integer inputs, including values between 0 and 1 . Syntax is plotDirichlet $\left(\mathrm{a}_{1}, \mathrm{a}_{2}, \mathrm{a}_{3}\right)$.
e) Write some code for a Bayesian learner observing a series of ternary data, using a Dirichlet conjugate prior. The structure should be something like this:

Set prior, i.e. parameters $a_{1}, a_{2}, a_{3}$ for Dirichlet
Define true generating probabilities for the observable, i.e. $q_{1}, q_{2}, q_{3}$ with $\Sigma q=1$
Loop through trials
Sample an observation according to the probabilities in $q$ [ 0 (i)=find (rand<cumsum (q),1)]
Update Dirichlet parameters for posterior
Plot posterior

